



(12) **EUROPEAN PATENT APPLICATION**

(43) Date of publication:
06.07.2016 Bulletin 2016/27

(51) Int Cl.:
G01M 5/00 (2006.01)

(21) Application number: **14200502.4**

(22) Date of filing: **29.12.2014**

(84) Designated Contracting States:
AL AT BE BG CH CY CZ DE DK EE ES FI FR GB GR HR HU IE IS IT LI LT LU LV MC MK MT NL NO PL PT RO RS SE SI SK SM TR
 Designated Extension States:
BA ME

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(54) **A method and a system for structural health monitoring**

(57) A method for detecting structural damage of a structure (110), the method comprising steps of: receiving measurement data (x_t, y_t) from sensors (111, 112) coupled to the structure (110); in a computing device (120), monitoring the measurement data (x_t, y_t) to detect

a change from homoscedasticity to heteroscedasticity; and upon detecting a change from homoscedasticity to heteroscedasticity, outputting via an output interface (124) a signal indicating a structural damage.

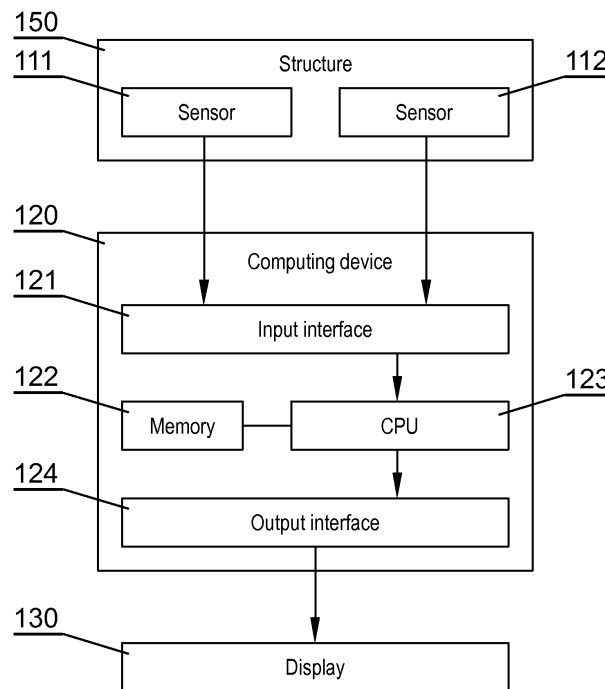


Fig. 1

Description

TECHNICAL FIELD

5 **[0001]** The present invention relates to a method and a system structural health monitoring.

BACKGROUND

10 **[0002]** Structural Health Monitoring (SHM) plays an important role in maintenance of air, sea and land transportation, civil infrastructure, energy supply systems and many industrial installations. Damage detection in SHM involves: observation of structure or mechanical systems over time using periodically spaced measurements, extraction of damage-sensitive features from these measurements and statistical analysis of these features, to determine the current operational condition.

15 **[0003]** One of the major problems that limits practical and commercial applications of SHM to real engineering structures is that many data and extracted features used for damage detection are often corrupted by ambient environment and operational conditions. A reliable SHM system should be able to distinguish between changes caused by damage and those caused by time-varying environmental and operational conditions.

[0004] Various approaches were developed to deal with the undesired effect of environmental and operational variability in data used for damage detection.

20 **[0005]** A US patent application US20080011091 discloses a method for measuring loading and temperature in structures and materials by measuring changes in natural frequencies, comprising: placing at least one sensor on the element, the at least one sensor being capable of sensing at least one of resonant vibrational frequency changes and phase angle changes indicative of changes in load on the element, the at least one sensor also measuring changes in temperature; applying a computer control unit to operate the at least one sensor so as to measure at least one of a change in a resonance frequency in the element and a change in phase angle sensed by the at least one sensor as a result of a change in load on the member; and applying the computer control unit to measure changes in temperature, and to apply a temperature calibration to calculate the load on the member.

25 **[0006]** A PCT patent application WO2014176625 discloses a method for determining a health condition of a structure, such as a bridge, based on vibration data measured of the bridge. Measured vibration data is calibrated. Features are then extracted from the calibrated data and a support vector machine classifier is then applied to the extracted features to determine the health condition of a part of the structure. Training of the support vector machine classifier by a machine learning process is also described.

30 **[0007]** It is well known that time series responses from engineering structure often exhibit nonlinear behavior. Moreover, operational and/or environmental common trends are typically believed to be nonlinearly related with response data used for damage detection. If this is the case then the linear cointegration theory may be no longer suitable for damage detection and a nonlinear approach to cointegration is needed.

35 **[0008]** Nonlinear cointegration has been studied in many different contexts. Research work demonstrates that non-stationarities and nonlinearities should be analysed simultaneously since in time series analysis nonlinearities often exist in a nonstationary context. However, it is not easy to reach this goal because the inherent difficulties in analyzing nonlinear time series models within a stationary and ergodic framework are enhanced in nonstationary contexts. This issue is also true for cointegration analysis when it is used for nonlinear and nonstationary processes. Hence, other definitions of stationarity and nonstationarity are needed in order to characterize better the usual notion of stationary $I(0)$ and nonstationary $I(1)$ time series, and cointegration in nonlinear contexts. The concepts of short memory and extended memory variables have been commonly used to ease this task.

40 **[0009]** A time series is said to be short memory if its information decays through time. In particular, a variable is short memory in mean (or in distribution) if the conditional mean (or conditional distribution) of the variable at time t given the information at time $t-h$ converges to a constant as h diverges to infinity. Shocks in short memory time series have transitory effects. In contrast, a time series is said to be extended memory in mean (or in distribution) if it is not short memory in mean (or in distribution). Shocks in extended memory time series have permanent effects. In heuristic terms, with respect to a "short memory in distribution" series, the influence of the past observations of the series decreases gradually. In the very far future, the past information is of no relevance to explain the stochastic patterns of the random variable. Whereas, regarding a "short memory in mean" series, the past information becomes progressively less relevant in the conditional h -step forecast in mean of the series. This means that the concept of short memory in this context can be considered as a somewhat stronger condition than stationarity; and the concept of extended memory can be thought

45 as a fairly stronger condition than nonstationarity.

50 **[0010]** The following general definition of nonlinear cointegration has been proposed (J.C. Escanciano, A. Escribano, Econometrics: Nonlinear Cointegration, Encyclopedia of Complexity and Systems Science, Berlin: Springer-Verlag (2009) 2757-2769): "If two or more series are of extended memory, but a nonlinear transformation of them is short

memory, then the series are said to be nonlinearly cointegrated". However, a simpler and more common definition of nonlinear cointegration is used in the current investigations. Two nonstationary time series x_t and y_t are nonlinearly cointegrated if there exists a nonlinear function f such that $z_t = f(x_t, y_t)$ is stationary.

5 **[0011]** This simplified definition is still quite general to be fully operative, and moreover, identification problems might be arisen in this general context. Hence, in practice some classes of function f are often used to avoid such identification problems. For example, one can consider a function of the form $z_t = g(x_t) - h(y_t)$, and estimate g and h by using nonparametric estimation procedures. Another approach is to consider transformations of the form, i.e. $z_t = f(x_t)+y_t$, or $z_t = f(x_t)-y_t$. The second approach is believed to be convenient for exploring nonlinear cointegration, therefore it has been used in this description. However, the question still remains how to construct a nonlinear function f .

10 **[0012]** Nonlinear cointegration has been recently proposed for SHM application (E.J. Cross, K. Worden, Approaches to nonlinear cointegration with a view towards applications in SHM, Journal of Physics: Conference Series 305 (2011) 012069 - referred to as Cross et al.). This section recalls one theoretical example from this work to demonstrate the major difficulty associated with nonlinear cointegration.

15 **[0013]** A system is described using two different time series variables x_t and y_t . The first variable x_t reacts linearly with respect to some external disturbance (D) whereas the second variable y_t reacts nonlinearly, in a quadratic way, to the same external disturbance. The given variables take the form

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$$x_t = \alpha T_t + \varepsilon_{1,t} \tag{1}$$

25
$$y_t = \beta T_t^2 + \varepsilon_{2,t} \tag{2}$$

where α, β are constants, T_t is some deterministic trend caused by the external disturbance, $\varepsilon_{1,t}$ and $\varepsilon_{2,t}$ are random normally-distributed processes. In this example, the driving trend (i.e. the external disturbance) is set to be linearly increasing with time t .

30 **[0014]** It is clear that a linear combination of both variables will not result in a stationary sequence. However, some combination of y and the square of x should produce a comparatively stationary signal

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$$z_t = f(x_t, a_1) + a_2 y_t = a_1 x_t^2 + a_2 y_t \tag{3}$$

[0015] In this case the nonlinear function $f(x_t, a_1) = a_1 x_t^2$ is referred to as the cointegrating function. If the parameters a_1 and a_2 can be found so that z_t is stationary, the vector $[a_1, a_2]$ will be analogous to the linear cointegrating vector, and x and y will be nonlinearly cointegrated. Two approaches to nonlinear cointegration have been proposed in to estimate the parameters a_1 and a_2 .

[0016] The first one is a variation of the well-established Johansen's linear cointegration procedure based on the use of an augmented basis. In this way, x_t^2 was first computed and then included as variable in its own right in the Johansen's

45 procedure. This means that the input to the Johansen's procedure would be $\{x_t^2, y_t\}$ and hence the only difference in

the approach to the Johansen's procedure is a manipulation of the form of the variables that are linearly combined. The second approach is a nonlinear optimisation routine based on differential evolution, where the objective is to choose parameters $[a_1, a_2]$ so that z_t is as stationary as possible. Two options were selected for the objective function: (1) to minimise the variance of the cointegration residuals; (2) to maximise the stationarity of the cointegration residuals. The results showed that nonlinear cointegrating vectors were created, the nonlinear trend was successfully removed, and stationary residuals were found for the times series by both proposed approaches. However, the variance of cointegration residuals was increasing with time, although cointegrated variables were mean stationary. This behavior - known as *heteroscedasticity* - implies that cointegration residuals are not strictly stationary. The problem can be further analysed when Eq. (3) is expanded using x_t and y_t from Eq. (1) and Eq. (2), respectively to give

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$$z_t = f(x_t, a_1) + a_2 y_t = a_1 x_t^2 + a_2 y_t = a_1 (\alpha T_t + \varepsilon_{1,t})^2 + a_2 (\beta T_t^2 + \varepsilon_{2,t}) \quad (4)$$

$$\Rightarrow z_t = a_1 (\alpha^2 T_t^2 + 2\alpha T_t \varepsilon_{1,t} + \varepsilon_{1,t}^2) + a_2 (\beta T_t^2 + \varepsilon_{2,t})$$

[0017] The differential evolution and the Johansen procedure selects the parameters a_1 and a_2 so that the quadratic deterministic trends cancel each other out. The cointegration residual from the nonlinear combination in Eq. (4) will then take the form

$$z_t = 2a_1 \alpha T_t \varepsilon_{1,t} + a_1 \varepsilon_{1,t}^2 + a_2 \varepsilon_{2,t} \quad (5)$$

[0018] The remaining terms in Eq. (5) include not only randomly distributed noise but also a term depending on $T_t \varepsilon_{1,t}$. This term is responsible for the increasing variance, because as mentioned above, the driving trend of the simulated time series in this case was chosen to be linearly increasing with time t , i.e. T_t is a linear deterministic trend in this case. Thus regardless of the nature of the driving trend, a combination of the type given in Eq. (3) will always result in a cointegration residual that has variance dependent on that trend. As a result heteroscedasticity will be always present in cointegration residuals obtained from the proposed approach. When the method is used for damage detection, this can lead to serious consequences.

[0019] It is clear that another solution is needed for an improved method and system for structural health monitoring, in order to create a stationary residual independent of the driving trends of the original time series.

SUMMARY

[0020] There is presented a method for detecting structural damage of a structure, the method comprising steps of: receiving measurement data (x_t, y_t) from sensors coupled to the structure; in a computing device, monitoring the measurement data (x_t, y_t) to detect a change from homoscedasticity to heteroscedasticity; and upon detecting a change from homoscedasticity to heteroscedasticity, outputting via an output interface a signal indicating a structural damage.

[0021] Preferably, the method further comprises: calculating nonlinear cointegration (z_t) for the measurement data (x_t, y_t) as $z_t = f(x_t) - y_t$; setting the condition that the nonlinear cointegration (z_t) becomes a zero mean series; transforming

the cointegration residual using a transformation R , wherein $R = \frac{1}{\sqrt{1 + [F'(x_t)]^2}}$ to obtain a transformed cointegration

residual as $z_t^R = \frac{z_t}{\sqrt{1 + [F'(x_t)]^2}}$; calculating a modified cointegration residual $z_t^* = \frac{f(x_t) - y_t}{\sqrt{1 + [f'(x_t)]^2}}$; and deter-

mining whether the modified cointegration residual z_t^* is homoscedastic or heteroscedastic.

[0022] Preferably, determining whether the modified cointegration residual z_t^* is homoscedastic or heteroscedastic is performed using the adapted Breusch-Pagan test procedure.

[0023] There is also presented a computer program comprising program code means for performing all the steps of the computer-implemented method as described herein when said program is run on a computer, as well as a computer readable medium storing computer-executable instructions performing all the steps of the computer-implemented method as presented herein when executed on a computer.

[0024] There is also presented a computer-implemented system for detecting damage in an object under test, the system comprising a computing device configured to perform the method presented herein.

BRIEF DESCRIPTION OF DRAWINGS

[0025] These and other objects of the invention presented herein are accomplished by providing a system and method for structural health monitoring. Further details and features of the present invention, its nature and various advantages will become more apparent from the following detailed description of the preferred embodiments shown in a drawing, in which:

Fig. 1 presents a general schematic of the system for structural health monitoring;
 Figs. 2-6 present experimental measurement results.

DETAILED DESCRIPTION

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 [0026] Fig. 1 presents a general schematic of the system 100 for structural health monitoring. The system comprises at least two sensors 111, 112 coupled to a structural element 150 to be monitored. Measurement data x_t, y_t from the sensors 111, 112 is communicated, in a wired or wireless manner, to a computing device 120 configured to execute the method of the present invention, as described below. The computing device 120 comprises an input interface 121 for receiving data from the sensors 111, 112. The computing device 120 may be controlled by one or more computer programs. Such computer programs are typically executed by utilizing the computing resources in the computing device. Applications are stored on a non-transitory medium. An example of a non-transitory medium is a non-volatile memory, for example a flash memory while an example of a volatile memory is RAM. The memories are jointly indicated by reference 122. The computer instructions are executed by a processor (CPU - Central Processing Unit) 123. Data calculated by the CPU are then output as a signal indicating structural damage via an output interface 124. The output interface 124 may be a data interface configured to provide the data in a predefined format, such as data to be used by other computing devices. Alternatively, the output interface 124 may be a graphical user interface configured to present various calculated data to the user in form of e.g. plots, on a display 130.

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 [0027] A homoscedastic nonlinear cointegration is proposed for use in the present method. The method overcomes the heteroscedastic problem related to cointegration residuals by offering a variance stabilizing nonlinear cointegration.

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 [0028] Following the work presented by Cross et al., two time series can be defined as

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$$x_t = T_t + \varepsilon_{1,t} \tag{6}$$

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$$y_t = F(T_t) + \varepsilon_{2,t} \tag{7}$$

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 where:

T_t is some deterministic trend caused by the external disturbance,
 $\varepsilon_{1,t}, \varepsilon_{2,t}$ are independent and identically distributed random variables and function $F(T_t)$ has a continuous and differentiable first derivative.

35
 It is assumed that $\varepsilon_{1,t}$ and $\varepsilon_{2,t}$ have zero mean and they are relatively small to T_t and independent from T_t . It is important to note that when the constants α and β are set to $\alpha = \beta = 1$, and $F(T_t) = T_t^2$, Eq. (6) and Eq. (7) are identical to Eq. (1) and Eq. (2), respectively. This means that the investigation presented herein has used the same time series variables as in Cross et al.

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 [0029] A simple form of nonlinear cointegration given as

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$$z_t = f(x_t) + y_t \tag{8}$$

can be proposed. This form can be rewritten - without loss of generality - into

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$$z_t = f(x_t) - y_t \tag{9}$$

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 due to the calculation procedure that will be more lucid for Eq. (9) than Eq. (8). When x_t and y_t from Eq. (6) and Eq. (7), respectively are substituted to Eq. (9) the cointegration residual can be obtained as

$$z_t = f(x_t) - y_t = f(T_t + \varepsilon_{1,t}) - F(T_t) - \varepsilon_{2,t} \quad (10)$$

5 **[0030]** It is clear that for z_t to become a zero mean series one can select the cointegrating function $f=F$, which thus yields

$$z_t = F(T_t + \varepsilon_{1,t}) - F(T_t) - \varepsilon_{2,t} \quad (11)$$

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Next, using the first degree Taylor approximation formula defined as

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$$F(x + a) \approx F(x) + F'(x)a + o(|a|^2) \quad (12)$$

for $F(T_t + \varepsilon_{1,t})$ results in

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$$F(T_t + \varepsilon_{1,t}) \approx F(T_t) + F'(T_t)\varepsilon_{1,t} + o(|\varepsilon_{1,t}|^2) \quad (13)$$

Then, substituting Eq. (13) into Eq. (11) yields

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$$z_t = F(T_t) + F'(T_t)\varepsilon_{1,t} + o(|\varepsilon_{1,t}|^2) - F(T_t) - \varepsilon_{2,t} \quad (14)$$

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The above equation can be approximated as

$$z_t \approx F'(T_t)\varepsilon_{1,t} - \varepsilon_{2,t} \quad (15)$$

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[0031] Eq. (15) shows that the cointegration residual z_t is zero mean, but its variance is not constant and strongly depends on the deterministic trend T_t . Since $\varepsilon_{2,t}$ is independent from $\varepsilon_{1,t}$ and T_t the variance of z_t can be estimated as

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$$Var(z_t) \approx Var(F'(T_t)\varepsilon_{1,t} - \varepsilon_{2,t}) = Var(F'(T_t)\varepsilon_{1,t}) + Var(-\varepsilon_{2,t}) \quad (16)$$

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$$\Rightarrow Var(z_t) \approx [F'(T_t)]^2 Var(\varepsilon_{1,t}) + Var(\varepsilon_{1,t}) = Var\left(\varepsilon_{1,t} \sqrt{1 + [F'(T_t)]^2}\right) \quad (17)$$

From Eq. (6) one can take that $x_t \approx T_t$ then Eq. (17) becomes

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$$Var(z_t) \approx Var\left(\varepsilon_{1,t} \sqrt{1 + [F'(x_t)]^2}\right) \quad (18)$$

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[0032] Eq. (15) shows that variance $Var(z_t)$ is not constant because it depends on $\sqrt{1 + [F'(x_t)]^2}$. This is where

the problem of heteroscedasticity appears. In order to solve this problem, the transformation R for the cointegration residual z_t can be proposed as

$$R = \frac{1}{\sqrt{1 + [F'(x_t)]^2}} \quad (19)$$

[0033] This complies well with the problem of the removal of the effect of heteroscedasticity. Finally, one obtains the transformed cointegration residual z_t^R that has the form

$$z_t^R = \frac{z_t}{\sqrt{1 + [F'(x_t)]^2}} \quad (20)$$

[0034] It is important to note that that the proposed solution is possible if and only if the cointegration residual z_t is zero mean. This requirement has been met in the case of the transformed cointegration residual proposed above. Clearly,

$\sqrt{1 + [F'(x_t)]^2}$ is constant if and only if $F'(x_t)$ is constant. Furthermore, when $F'(x_t)$ is constant then $F(x_t)$ is linear, which thus implies that x_t and y_t are linearly related. This explains why cointegration residuals - created in the context of linear cointegration - are homoscedastic without any modification.

[0035] When Eq. (9) is met together with the condition for z_t to become a zero mean series (i.e. when $f = F$) then Eq. (20) becomes

$$z_t^* = \frac{f(x_t) - y_t}{\sqrt{1 + [f'(x_t)]^2}} \quad (21)$$

[0036] This equation gives the *modified cointegration residual* z_t^* that is not only zero mean but also homoscedastic. It is important to note that the solution proposed by Eq. (21) is general and thus can be applied for any form of unknown functions (trends).

[0037] In order to test for the presence of heteroscedasticity in the analysed data, various tests can be used. In particular, it is preferable to use an adapted Breusch-Pagan test procedure, as described in the publication "A simple test for heteroscedasticity and random coefficient variation" by T.S. Breusch, A.R. Pagan, *Econometrica* 47 (1979) 1287-1294. In principle, the Breusch-Pagan test tests for conditional heteroscedasticity, i.e. it tests whether the estimated variance of the residuals from a regression are dependent on the values of the independent variables. The approach is based on the Lagrange Multiplier (LM) test statistic with an assumption that the error terms are normally distributed.

[0038] The linear time series regression model for one independent variable x_{1t} can be written as

$$y_t = \alpha_0 + \alpha_1 x_{1t} + \varepsilon_t \quad (22)$$

where y_t is a dependent variable, ε_t is a random error term (or a residual), and (α_0, α_1) are coefficients.

[0039] Next, in order to test for the presence of heteroscedasticity in the residual ε_t the auxiliary regression model is formed as

$$\varepsilon_t^2 = \beta_0 + \beta_1 x_{1t} + error \quad (23)$$

where (β_0, β_1) are coefficients.

[0040] The Breusch-Pagan heteroscedasticity test is performed by regressing the squared residuals directly on the independent variables. In the linear time series regression model, one can assume that the mean of the residual ε_t is

5 zero. Hence, the estimated variance of the residual (i.e. ε_t^2) in Eq. (23) is constant if and only if it is independent from the independent variable x_{1t} . If this is the case then β_1 should be close/equal to zero. The LM test statistic is used to evaluate the significance of β_1 .

[0041] It should be noted that only one independent variable has been used in the current investigations. In general case when more than one independent variables are employed then the test statistic equals NR^2 , where N is the sample size and R^2 is the coefficient of determination in the auxiliary regression.

[0042] Since the original Breusch-Pagan test is only used to test for heteroscedasticity in a linear regression model. Hence, the test has been adapted to be able to use in this paper, i.e. to test for the presence of heteroscedasticity in the cointegration residuals obtained from nonlinear cointegration analysis, which is explained in the following.

[0043] Firstly, the linear regression model in Eq. (22) can be rewritten to the form $y_t - f(x_t) = \varepsilon_t$, where $f(x_t) = \alpha_0 + \alpha_1 x_{1t}$. In this form, the concepts of independent and dependent variables become invisible.

[0044] Secondly, in general case since the mean of ε_t should not be assumed to be equal to zero so that the estimated variance of the residual ε_t may take the form $[\varepsilon_t - E(\varepsilon_t)]^2$, where $E(\varepsilon_t)$ is the mean of ε_t . Then, the auxiliary regression model can be formed as

$$[\varepsilon_t - E(\varepsilon_t)]^2 = \beta_0 + \beta_1 y_t + error \quad (24)$$

[0045] Following the same discussion as above, Eq. (24) shows that the residual ε_t is homoscedastic if the term on the left, i.e. $[\varepsilon_t - E(\varepsilon_t)]^2$, is independent from y_t . This implies that the coefficient β_1 should be equal to zero. In the current investigations, the significance of β_1 is assessed by using the Student's t-test statistic, which is shortly described as follow: The null hypothesis is true, i.e. there is no heteroscedasticity in the residual ε_t , if the coefficient β_1 is insignificant (i.e. $p\text{-value} > 0.05$). Conversely, the alternative is true, i.e. there is heteroscedasticity in the residual ε_t , if the coefficient β_1 is significant (i.e. $p\text{-value} < 0.05$).

[0046] It should be noted that x_{1t} can be considered in the auxiliary regression model in Eq. (24), instead of y_t . Since y_t has been considered in the auxiliary regression model, the correlation between the absolute value of ε_t and x_{1t} can be checked to determine how the values of x_{1t} deviate from ε_t .

35 EXAMPLES

[0047] Examples that explain the homoscedastic nonlinear cointegration method and illustrate its application to nonlinear trend removal and possible damage detection are presented in this section. Three different time series - i.e. one with a nonlinear quadratic deterministic trend, simulated vibration data and experimental wind turbine data - are analysed.

40 Quadratic cointegrating function

[0048] This section recalls the nonlinear cointegrating function $f(x_t) = x_t^2$ that has been used in Cross et al. The objective is to demonstrate that the homoscedastic nonlinear cointegration method can remove heteroscedasticity from cointegration residuals.

[0049] When the nonlinear cointegration form given by Eq. (9) is used the cointegration residual can be calculated as

$$z_t = f(x_t) - y_t = x_t^2 - y_t \quad (25)$$

[0050] Similarly, the homoscedastic nonlinear cointegration - given by Eq. (21) can be used. This results in the modified cointegration residual that takes the form

$$z_t^* = \frac{f(x_t) - y_t}{\sqrt{1 + [f'(x_t)]^2}} = \frac{x_t^2 - y_t}{\sqrt{1 + [2x_t]^2}} \quad (26)$$

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[0051] Fig. 2a and Fig. 2b present the cointegration residual z_t and the modified cointegration residual z_t^* , respectively. The results show that the nonlinear quadratic deterministic trend was successfully removed in both cases. However, the variance of z_t increases with time (i.e. z_t is heteroscedastic), whereas the variance of z_t^* is relatively stable (i.e. z_t^* is homoscedastic). The adapted Breusch-Pagan test procedure was used to confirm these results. Consequently, the test statistic for z_t is significant because $p\text{-value} < 0.0001 < 0.05$, therefore z_t is heteroscedastic. In contrast, the relevant test statistic for z_t^* is insignificant because $p\text{-value} = 0.318 > 0.05$, so that z_t^* is homoscedastic. This simple example demonstrates that the proposed homoscedastic nonlinear cointegration method can remove heteroscedasticity from cointegration residuals.

20 Simulated bridge data

[0052] The example presented in this section shows that the proposed homoscedastic nonlinear cointegration method can remove undesired nonlinear trends from data and reliably detect structural damage. Simulated data are used in this example.

25 **[0053]** Vibration/modal data are often used to monitor civil engineering structures. Various methods have been developed in practice for structural damage detection. Often natural frequencies are used to obtain information about possible damage. The assumption is structural damage reduces stiffness leading to the reduction of natural frequencies. However, it is well known that natural frequencies of large civil engineering structures - such as bridges - can be affected by environmental conditions (e.g. temperature or heavy rain).

30 **[0054]** The nonlinear relation between modal properties of the bridge and daily temperature was explored in this example. The nonlinear dependence of the first natural frequency (given in Hz) of the bridge on temperature (given in °C) was simulated for 5760 measurements using the $f(t)$ function that was defined as

$$f(t) = \begin{cases} 4.4 & \text{for } t \leq -2 \\ 4.2 + 0.2 \cos[(t+2)\pi] & \text{for } t \in (-1, -2) \\ 4.0 & \text{for } t \geq -1 \end{cases} \quad (27)$$

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[0055] Once the data were simulated using Eq. (27), a random number from a normal distribution noise process with the mean value $\mu = 0$ and the standard deviation $\sigma = 0.02$ - was added to each simulated value $f(t)$ to reflect measurement noise. The natural frequency was simulated for daily temperature variations from the minimum value around -4 °C (night time) to the maximum value around 6 °C (day time). The $f(t)$ function was defined to exhibit an approximate shape of saturation nonlinearity - i.e. the frequency-temperature relationship exhibited a linear trend for the values of temperature from -2 °C to -1 °C; below and above these temperatures the natural frequency was approximately 4.4 and 4.0 Hz, respectively. Fig. 3 gives the simulated data demonstrating how the first natural frequency of the bridge changes nonlinearly with temperature. Over one day and night of monitoring, the most dramatic change of natural frequency can be observed around -1.5 °C.

55 **[0056]** It is important to note that the natural frequency data were simulated to resemble the experimental vibration data from the Z24 bridge - a pre-stressed concrete highway bridge - that joins two small villages in Switzerland between Bern and Zurich (B. Peeters, G.D. Roeck, One year monitoring of the Z24-Bridge: environmental influences versus damage events, Earthquake Engineering & Structural Dynamics 30 (2001) 149-171.). Although, the original experimental data were used in the nonlinear cointegration example given by Cross et al., the simulated data were used intentionally in the current investigations to account for three different structural conditions of the bridge, as described below. The effect of simulated damage was simply introduced to the first natural frequency in a nonlinear manner by multiplying the

$f(t)$ function by the factor \sqrt{k} , where k represents the stiffness of the analysed bridge. It is well known that the change of stiffness alters the natural frequency. Thus in this example the analysed natural frequency was modified by both effects, i.e. environmental condition (temperature) and structural damage. The condition of the bridge (or in other words the severity of damage) was simulated following three steps:

1. The bridge was undamaged for the first two days, i.e. the value of k was kept 1, and therefore the first natural frequency depended only on the daily temperature variations.
2. During the third day, the bridge stiffness was reduced gradually and finally reached 95% of the normal value by the end of the day (i.e. $k = 0.95$); this simulates the initial structural deterioration (or damage). Hence, since the beginning of the third day the first natural frequency depended on both, i.e. temperature and damage.
3. At the beginning of the fourth day, the bridge stiffness was reduced suddenly from 95% to 90% of the normal stiffness value (i.e. $k = 0.9$); the deteriorating damage condition reached the maximum severity.

[0057] Once the data were simulated to reflect four days of monitoring, the nonlinear cointegration procedure and the homoscedastic nonlinear cointegration procedure were used to obtain the original cointegration residual z_t and the modified cointegration residual z_t^* , respectively.

[0058] Fig. 4a and Fig. 4b present the results obtained for the z_t and z_t^* , respectively. The first and second dash-dotted vertical line in these figures indicates the start of the third day (or deteriorating continuous damage) and the start of the fourth day (or the moment of sudden damage), respectively. When temperature variations are present in the data but the bridge is not damaged (the first two days) the original cointegration residual z_t in Fig. 4a exhibits four (two per day) noisy patterns of local peaks corresponding to daily temperature variations (following the characteristic given in Fig. 3). In contrast, the modified cointegration residual z_t^* in Fig. 4b is much smoother and does not display any clear noisy patterns of peaks. This implies that the variations of temperature manifest themselves strongly in the case of the original cointegration residual z_t , but much less in the case of the modified cointegration residual z_t^* . In other words, the nonlinear temperature-dependent trend was significantly removed from the z_t^* , but not completely removed from the z_t . The adapted Breusch-Pagan test procedure was used to test for the presence of heteroscedasticity in both cointegration residuals. The test statistic for z_t was significant because $p\text{-value} < 0.0001 < 0.05$, indicating the heteroscedastic behaviour. In contrast, the test statistic for z_t^* was insignificant because $p\text{-value} = 0.213 > 0.05$. This statistic indicates that the nonlinear cointegration residual z_t^* is homoscedastic, as expected.

[0059] When temperature variations are present and the bridge has some initial progressing damage (the third day) the nonlinear temperature-dependent trend is again removed from the z_t^* , but not removed from z_t ; two noisy patterns of local peaks can be observed between 48 and 72 hours. In addition, there is an upward linear trend in that region for both characteristics in Fig. 4a and Fig. 4b. This trend reflects the gradual reduction of stiffness in the third day. However, by the end of the third day when the stiffness reduction reaches nearly 95% of the initial value, a significant change in the z_t^* characteristics can be observed in Fig. 4b. This change is a clear indication of damage. The adapted Breusch-Pagan test procedure was applied again to the third day data to test for the presence of heteroscedasticity in the z_t and z_t^* residuals. The test statistics for both residuals were significant because $p\text{-value} < 0.0001 < 0.05$, indicating that the z_t and z_t^* residuals are heteroscedastic.

[0060] During the fourth day - when temperature variations are present and the progressing damage of the bridge reaches the maximum level - Fig. 4a shows that the effect of temperature variations is still present in the z_t residual. However, a small change in the mean value of z_t can be observed at the second dash-dotted vertical line. This change looks like a "jump" in the analysed characteristics. This feature - also observed in the z_t^* residual in Fig. 4b is due to the abrupt reduction of stiffness from 95 to 90% of the initial stiffness value. Again, the adapted Breusch-Pagan test procedure was used to test for the presence of heteroscedasticity in both cointegration residuals and in this case the

results obtained are the same as in the third day, i.e. both statistics indicate that the z_t and z_t^* cointegration residuals are heteroscedastic.

[0061] In summary, the effect of temperature variations was not removed from the original nonlinear cointegration residual z_t . This residual is heteroscedastic in the entire analysed period of time. Consequently, this behavior can hide the effect of damage, thus leading to false-positive and/or false-negative damage detection scenarios. In contrast, the modified nonlinear cointegration residual z_t^* does not exhibit any temperature variations; the effect of temperature variations was successfully removed. As a result the z_t^* characteristics is homoscedastic until the monitored structure is damaged. Therefore the adapted Breusch-Pagan heteroscedasticity test can be used to detect damage initiation.

Wind turbine data

[0062] The third example illustrates the application of the homoscedastic nonlinear cointegration to the analysis of wind turbine operational data. Wind turbines are an important class of machines with growing penetration in the wind generation market.

[0063] Economics of the wind farms operation is very different from conventional sources. The main reason is the distribution of costs during the lifetime of the turbine. In the case of wind power generation, the capital expenditures are higher than for coal or gas plants (on the per kilowatt basis), but operational expenditures are much smaller, since there is no fuel cost. The main factor, which can ruin this financial model is unexpected failure, which results in costly repair and often months of machine unavailability. Hence, condition monitoring of wind turbines - in particular at the early stage of a fault occurrence - is an essential problem in wind turbines engineering.

[0064] Wind turbine data used in this paper originate from a series of experimental measurements for a wind turbine drivetrain with a nominal power of 2 MW. The SHM system for the wind turbine measured few operational parameters - i.e. wind speed, rotor speed, generator shaft rotational speed, generator temperature, gearbox temperature, generator voltage, generator current, active power, and active power delivered - for different operating conditions.

[0065] The relation between rotor speed (in *rpm*) and active power delivered (in *kW*) is important for wind turbine monitoring. This characteristic - for the rotor speed higher than 800 *rpm* - was used for the nonlinear cointegration analysis in this example. It is important to note that only the data from an intact wind turbine was available and used in these investigations. The objective was to demonstrate that the proposed method leads to the homoscedastic cointegration residuals for undamaged wind turbine, regardless of changes in the relation between rotor speed and active power delivered. The analysed data - shown in Fig. 5 - clearly display a nonlinear (or to be more precise bi-linear) trend.

[0066] The homoscedastic nonlinear cointegration method was applied to the data. The original cointegration residual z_t was computed using Eq. (9), whereas the modified cointegration residual z_t^* was calculated using Eq. (21). In this example, the rotor speed is the variable x_t , the active power is the variable y_t and the function $f(x)$ describes the relation between the rotor speed and the active power delivered.

[0067] Since $f(x_t)$ is unknown a local regression algorithm was used to estimate this function. An estimate of the function f at x_0 - i.e. $f(x_0)$ - was calculated using the nearest twelve values of x_0 from the training data; only the first 40% of the data were used for training to avoid overfitting. Then, the least squares regression was employed for this subset of values to estimate or predict the value of $f(x_0)$. The gradient (or slope) of the calculated regression was considered as the first derivative of the function f at x_0 , i.e. $f'(x)$. The entire procedure was repeated for all remaining values, i.e. x_1, x_2, \dots , which resulted in the estimate of the function f .

[0068] Fig. 6a and Fig. 6b give the original z_t and modified z_t^* nonlinear cointegration residuals, respectively. The first observation of the results indicates much larger amplitude variability (amplitude ranges from -40 to 40) for z_t if compared with z_t^* (amplitude variability ranges from -20 to 20). The adapted Breusch-Pagan test procedure was used to test for the presence of heteroscedasticity in both cointegration residuals. As a result, the test statistic for z_t was significant because $p\text{-value} < 0.0001 < 0.05$, indicating that z_t is heteroscedastic. In contrast, the test statistic for z_t^* was insignificant because $p\text{-value} = 0.168 > 0.05$, indicating that z_t^* is homoscedastic. Moreover, the correlations between the rotor speed (x_t) and the absolute values of z_t and z_t^* were calculated as 0.46 and 0.183, respectively. This means that the rotor speed x_t contains less information about deviation of z_t^* in comparison with z_t .

[0069] In summary, when the data from an intact wind turbine was analysed, the modified nonlinear cointegration residual z_t^* was homoscedastic for the whole data set investigated, regardless of changes in the relation between rotor speed and active power delivered. The homoscedasticity is a sign of undamaged condition. It is expected that the homoscedastic characteristics would be broken if the data for damage wind turbine were available and analysed.

Claims

1. A method for detecting structural damage of a structure (110), the method comprising steps of:

- receiving measurement data (x_t, y_t) from sensors (111, 112) coupled to the structure (110);
- in a computing device (120), monitoring the measurement data (x_t, y_t) to detect a change from homoscedasticity to heteroscedasticity; and
- upon detecting a change from homoscedasticity to heteroscedasticity, outputting via an output interface (124) a signal indicating a structural damage.

2. The method according to claim 1, further comprising:

- calculating nonlinear cointegration (z_t) for the measurement data (x_t, y_t) as $z_t = f(x_t) - y_t$;
- setting the condition that the nonlinear cointegration (z_t) becomes a zero mean series;
- transforming the cointegration residual using a transformation R, wherein $R = \frac{1}{\sqrt{1 + [F'(x_t)]^2}}$ to obtain a

transformed cointegration residual as $z_t^R = \frac{z_t}{\sqrt{1 + [F'(x_t)]^2}}$;

- calculating a modified cointegration residual $z_t^* = \frac{f(x_t) - y_t}{\sqrt{1 + [f'(x_t)]^2}}$; and

- determining whether the modified cointegration residual z_t^* is homoscedastic or heteroscedastic.

3. The method according to claim 2, wherein determining whether the modified cointegration residual z_t^* is homoscedastic or heteroscedastic is performed using the adapted Breusch-Pagan test procedure.

4. A computer program comprising program code means for performing all the steps of the computer-implemented method according to claim 1 or 2 or 3 when said program is run on a computer.

5. A computer readable medium storing computer-executable instructions performing all the steps of the computer-implemented method according to claim 1 or 2 or 3 when executed on a computer.

6. A computer-implemented system for detecting damage in an object under test, the system comprising a computing device (120) configured to perform the method of claims 1 or 2 or 3.

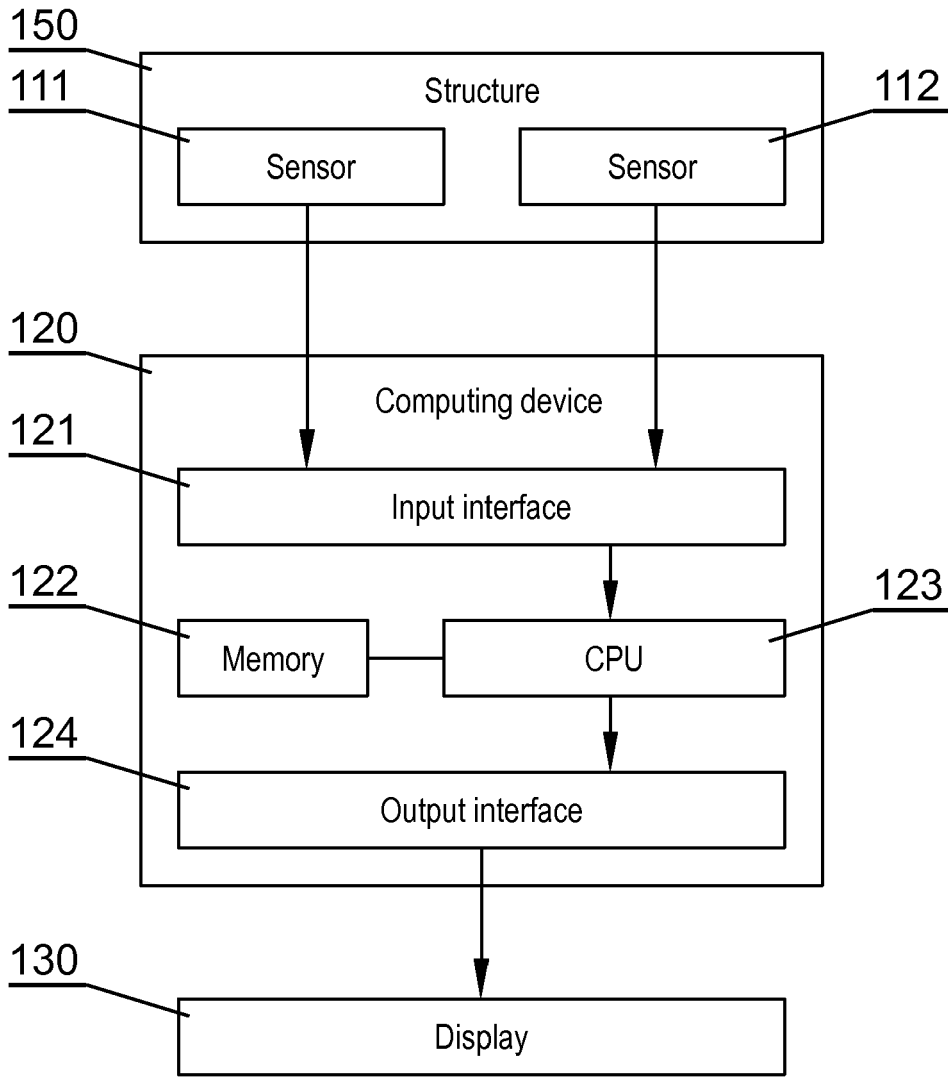


Fig. 1

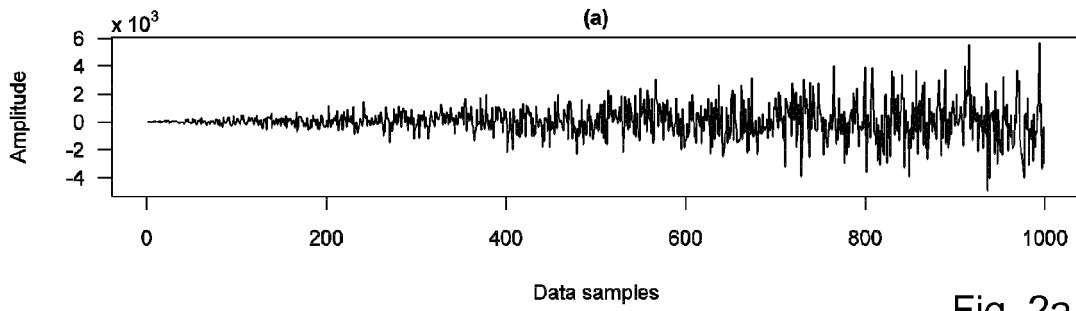


Fig. 2a

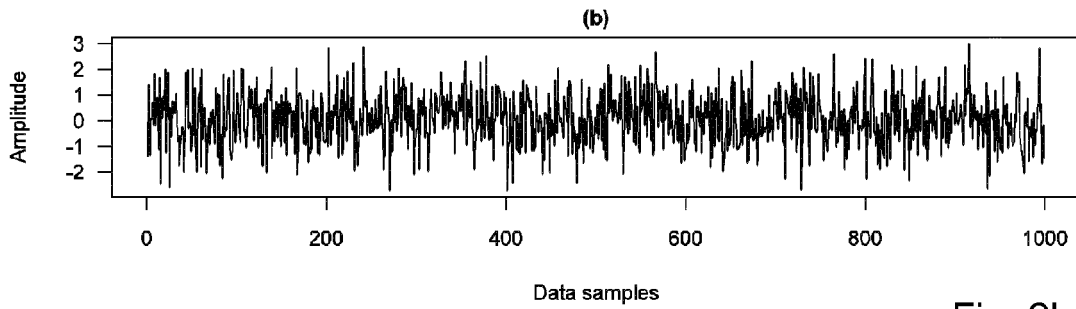


Fig. 2b

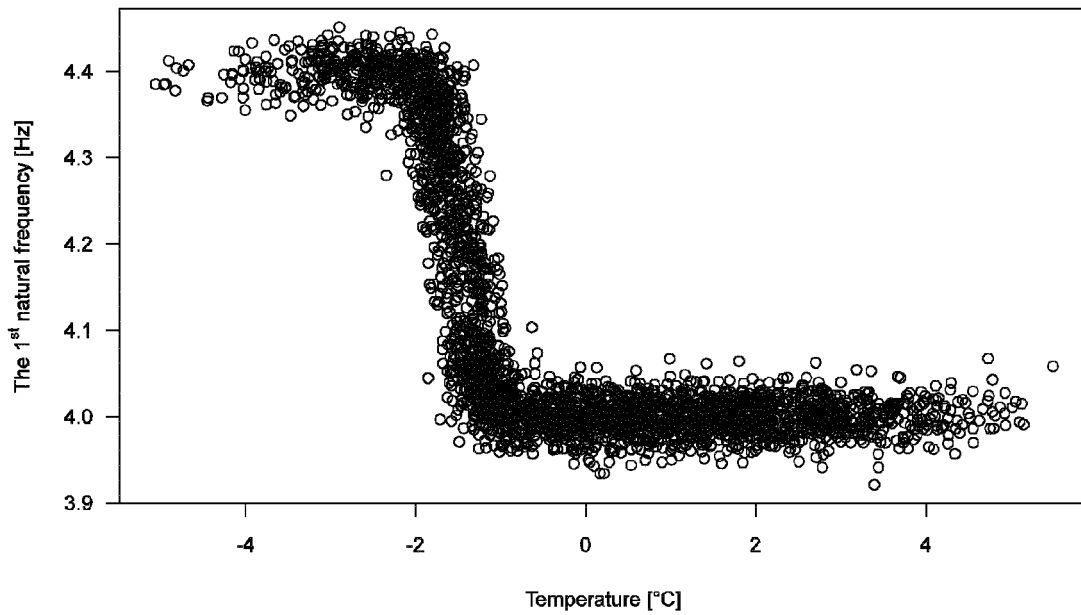
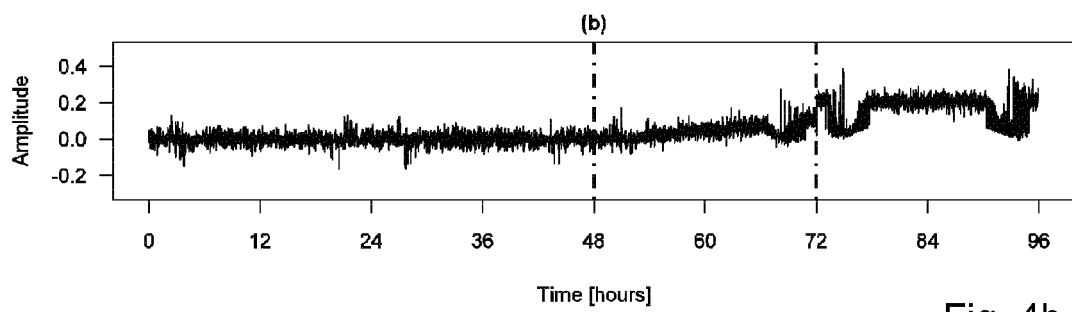
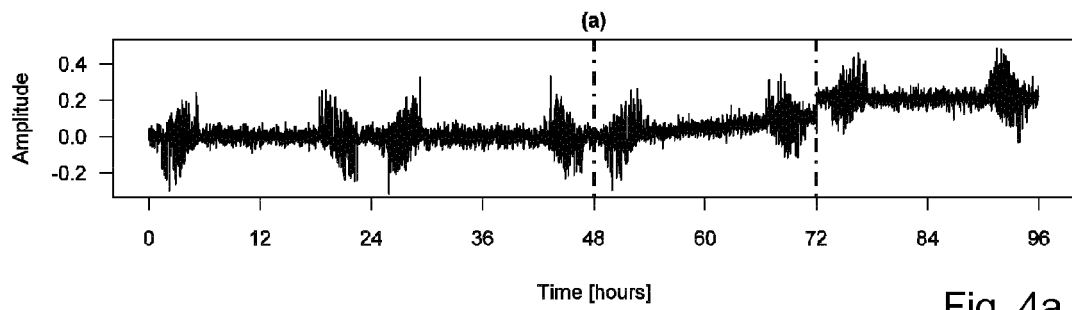


Fig. 3



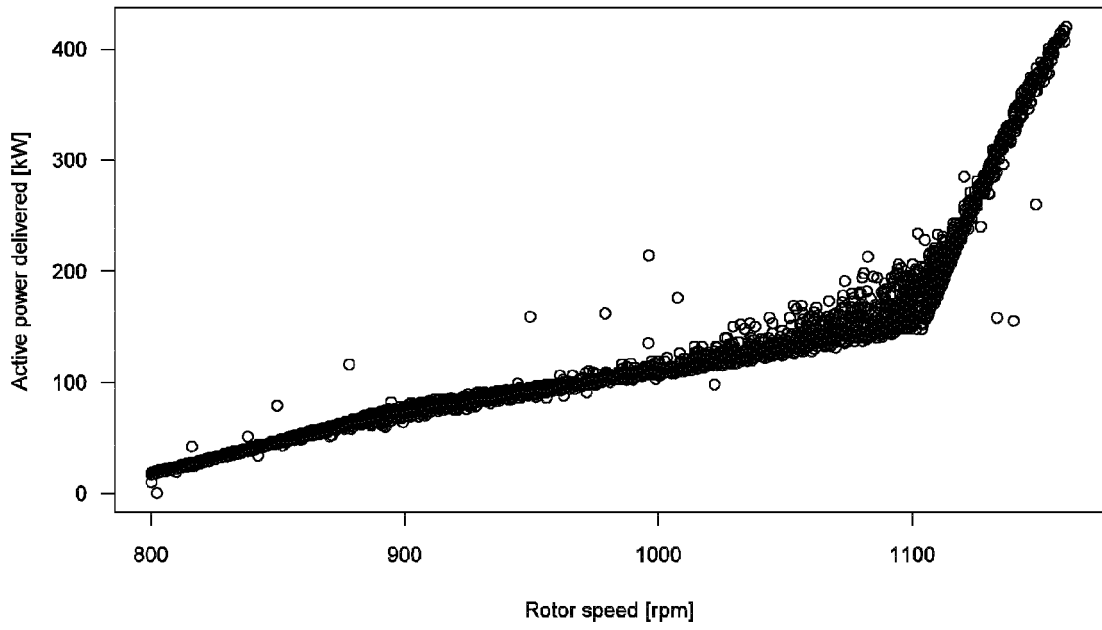


Fig. 5

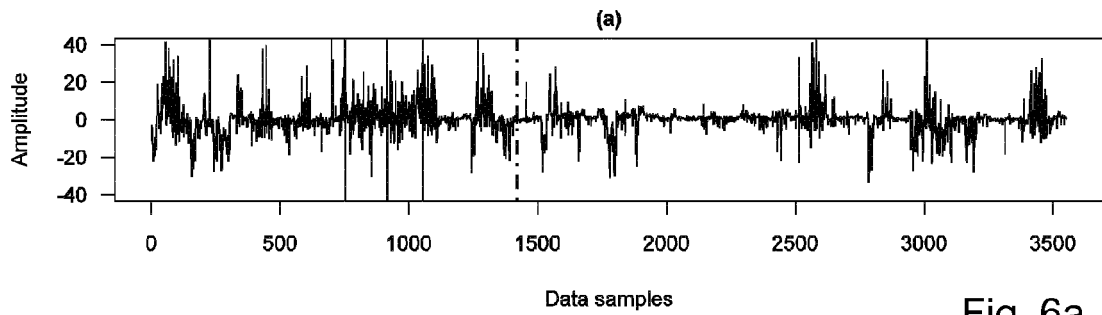


Fig. 6a

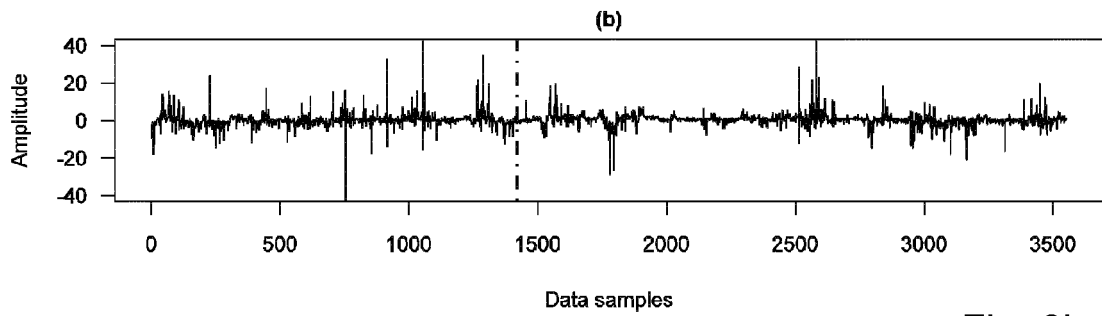


Fig. 6b



EUROPEAN SEARCH REPORT

Application Number
EP 14 20 0502

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DOCUMENTS CONSIDERED TO BE RELEVANT			
Category	Citation of document with indication, where appropriate, of relevant passages	Relevant to claim	CLASSIFICATION OF THE APPLICATION (IPC)
X	FOTIS P KOPSAFTOPOULOS ET AL: "Identification of Stochastic Systems Under Multiple Operating Conditions: The Vector Dependent FP-ARX Parametrization", CONTROL AND AUTOMATION, 2006. 14TH MEDITERRANEAN CONFERENCE ON, IEEE, PI, 1 June 2006 (2006-06-01), pages 1-6, XP031057328, DOI: 10.1109/MED.2006.236828 ISBN: 978-0-9786720-1-0	1,3-6	INV. G01M5/00
A	* the whole document * -----	2	
			TECHNICAL FIELDS SEARCHED (IPC)
			G01M G06F G01N
The present search report has been drawn up for all claims			
Place of search The Hague		Date of completion of the search 23 June 2015	Examiner Vytlačilová, Lenka
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